

DEDICATED TO PROFESSOR DOBIESLAW BOBROWSKI
ON THE OCCASION OF HIS 70th BIRTHDAY.

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GEOMETRIC APPROXIMATION OF MIXTURE OF DISCRETE LIFE TIME DISTRIBUTIONS

ABSTRACT. We consider a mixture of a distribution from the class (D)NBUE and a distribution, which describes the life time of small part of lower reliability elements. Given the first two moments of this mixture, the distribution can be approximated by a geometric distribution with some parametr p . An error of such approximation is also estimated.

KEY WORDS. mixture of discrete distributions, geometric distribution, error of approximation.

In two previous papers ([1], [2]) the possibility of approximation of unknown life time discrete distribution by the geometric distribution was studied. It was supposed that the unknown distribution belongs to the class (D)NBUE (discrete new better than used in expectation). In this paper we consider the same problem in a case when the unknown distribution is a mixture of a distribution from a given class (D)NBUE and other distribution. The second distribution describes the life time of rather small part of lower reliability elements e.g. elements stopped working on the beginning of exploitation time.

Let $(p_n) = (p_n, n \in N)$ be a mixture of independent discrete probability distributions (p'_n) and (p''_n) i.e.

$$(1) \quad p_n = (1 - \delta)p'_n + \delta p''_n, \quad n \in N,$$

where $\delta \in (0, 1)$. We will assume that the mixture (p_n) has a finite first and second moment $\mu > 1$, $\mu_2 > 1$, respectively.

In this paper (p'_n) and (p''_n) stands for the life time distribution of high reliability elements and faulty elements fraction, respectively. For the mixture we introduce the following notation for all $n \geq 0$:

$$R_n = \sum_{k=n+1}^{\infty} p_k,$$

and

$$\bar{G}_n = \frac{1}{\mu} \sum_{k=n}^{\infty} R_k.$$

R_n is call residual probability.

Clearly

$$R_n = (1 - \delta)R'_n + \delta R''_n,$$

where

$$R'_n = \sum_{k=n+1}^{\infty} p'_k \quad \text{and} \quad R''_n = \sum_{k=n+1}^{\infty} p''_k$$

is a residual probability for high reliability elements fraction and faulty elements fraction, respectively.

Similarly denoting by μ' the expected value of life time of good elements fraction and μ'' the expected value of life time of faulty elements fraction we have

$$\mu = (1 - \delta)\mu' + \delta\mu''.$$

For considered mixture we introduce following assumption for all $n \geq 0$

$$(2) \quad R_n \leq R'_n.$$

Under our assumption that fraction of high reliability elements belongs to (D)NBUE class so from the definition of such a class the inequality

$$(3) \quad \bar{G}'_n = \frac{1}{\mu'} \sum_{k=n}^{\infty} R'_k \leq R'_n$$

holds.

From the definition and assumptions (2) and (3) for our mixture we get following condition

$$\frac{1}{\mu} \sum_{k=n}^{\infty} R_k \leq \frac{1}{(1 - \delta)^2} R_n.$$

$$\text{Let} \quad \varepsilon = \varepsilon(\delta) = \frac{\delta(2 - \delta)}{(1 - \delta)^2} > 0 \quad \text{for} \quad \delta \in (0, 1),$$

and let

$$\Delta_n = (1 + \varepsilon)R'_n - \bar{G}'_n.$$

Then clearly

$$\bar{G}'_n \leq (1 + \varepsilon)R'_n,$$

and

- (i) $\Delta_n \geq 0,$
- (ii) $\Delta_0 = \varepsilon > 0,$
- (iii) $\lim_{n \rightarrow \infty} \Delta_n = 0.$

We begin with the following lemma

Lemma 1. Let the distribution (p_n) be a mixture of distributions defined by (1). In this case

$$(4) \quad \sum_{n=0}^{\infty} \Delta_n = \frac{\mu}{(1-\delta)^2} - \frac{\mu_2}{2\mu} - \frac{1}{2} =: \alpha_\delta$$

Proof. We have

$$\begin{aligned} \sum_{n=0}^{\infty} \Delta_n &= (1+\varepsilon) \sum_{n=0}^{\infty} R_n - \frac{1}{\mu} \sum_{n=0}^{\infty} \sum_{k=n}^{\infty} R_k = (1+\varepsilon) \sum_{n=0}^{\infty} R_n - \frac{1}{\mu} \sum_{k=0}^{\infty} (k+1)R_k = \\ &= (1+\varepsilon)\mu - \frac{1}{\mu} \sum_{k=0}^{\infty} kR_k - 1 = (1+\varepsilon)\mu - 1 - \frac{1}{2\mu} \sum_{n=0}^{\infty} (n^2 - n)p_n = \\ &= (1+\varepsilon)\mu - 1 - \frac{1}{2\mu} (\mu_2 - \mu) = \frac{\mu}{(1-\delta)^2} - \frac{\mu_2}{2\mu} - \frac{1}{2} \end{aligned}$$

and the proof is completed.

Our next result is the following technical lemma about the residual probability R_n .

Lemma 2. If (p_n) is a mixture of distributions defined by (1) then

$$(5) \quad R_n = \frac{1}{1+\varepsilon} q^n \left[1 - \frac{p}{q} \sum_{k=0}^{n-1} q^{-k} \Delta_k \right] + \frac{1}{1+\varepsilon} \Delta_n, \quad n \geq 0,$$

where

$$(6) \quad p = \frac{1}{(1+\varepsilon)\mu}, \quad \text{and} \quad q = 1 - p.$$

Proof. By the definitions

$$\Delta_n = [(1+\varepsilon)\mu - 1] \bar{G}_n - (1+\varepsilon)\mu \bar{G}_{n+1}.$$

Therefore

$$\bar{G}_{n+1} = q \bar{G}_n - p \Delta_n.$$

From this recurrence relation we obtain

$$\bar{G}_n = q^n \left[1 - \frac{p}{q} \sum_{k=0}^{n-1} q^{-k} \Delta_k \right].$$

Consequently

$$\bar{G}_n - \bar{G}_{n+1} = q^n p - q^{n-1} p^2 \sum_{k=0}^{n-1} q^{-k} \Delta_k + p \Delta_n.$$

Hence

$$\frac{1}{\mu} R_n = q^n p \left[1 - \frac{p}{q} \sum_{k=0}^{n-1} q^{-k} \Delta_k \right] + p \Delta_n$$

which implies (5).

The proof of the next theorem relies on the following formula for $R_n - q^n$.

Lemma 3. If (p_n) is a mixture of distributions defined by (1) then

$$(7) \quad (1 + \varepsilon)(R_n - q^n) = \sum_{k=1}^n q^{n-k} (\Delta_k - \Delta_{k-1}).$$

Proof. According to well known Abel's transformation (see e.g. [3])

$$\begin{aligned} \sum_{k=0}^{n-1} \Delta_k q^{-k} &= \Delta_n \sum_{k=0}^{n-1} q^{-k} - \sum_{k=0}^{n-1} \left[(\Delta_{k+1} - \Delta_k) \sum_{r=0}^k q^{-r} \right] = \\ &= \Delta_n \frac{1 - q^{-n}}{1 - q^{-1}} - \sum_{k=1}^n (\Delta_k - \Delta_{k-1}) \frac{1 - q^{-k}}{1 - q^{-1}} = \\ &= -\frac{q}{p} (1 - q^n) \Delta_n + \frac{q}{p} \sum_{k=1}^n (\Delta_k - \Delta_{k-1}) (1 - q^{-k}). \end{aligned}$$

Hence

$$\begin{aligned} \frac{p}{q} \sum_{k=0}^{n-1} \Delta_k q^{-k} &= -(1 - q^{-n}) \Delta_n + (\Delta_n - \Delta_0) - \sum_{k=1}^n q^{-k} (\Delta_k - \Delta_{k-1}) = \\ &= q^{-n} \Delta_n - \Delta_0 + \sum_{k=1}^n q^{-k} (\Delta_k - \Delta_{k-1}). \end{aligned}$$

Therefore by Lemma 2 we get (7).

Now we are ready to prove the main result of this paper.

Theorem. Let the discrete distribution (p_n) be a mixture of distributions defined by (1). Then there exists $q = 1 - \frac{1}{(1+\varepsilon)\mu} \in (0, 1)$ such that we get the following estimation

$$(8) \quad |R_n - q^n| \leq \beta_\delta,$$

where

$$(9) \quad \beta_\delta = (1 - \delta)^2 \alpha_\delta,$$

and α_δ defined by Lemma 1.

Proof. From Lemma 3 we get

$$R_n - q^n \leq \frac{1}{1+\varepsilon} \sum_{k=1}^n q^{n-k} \Delta_k \leq \frac{1}{1+\varepsilon} \sum_{k=0}^{\infty} \Delta_k = \frac{\alpha_\delta}{1+\varepsilon} = \beta_\delta.$$

On the other hand

$$\begin{aligned} R_n - q^n &\geq \frac{1}{1+\varepsilon} \sum_{k=1}^n q^{n-k} \Delta_{k-1} \geq -\frac{1}{1+\varepsilon} \sum_{k=1}^n \Delta_{k-1} \geq \\ &\geq -\frac{1}{1+\varepsilon} \sum_{k=0}^{\infty} \Delta_k = -\frac{\alpha_\delta}{1+\varepsilon} = -\beta_\delta. \end{aligned}$$

The theorem is proved.

Remark. If $\varepsilon = 0$ then we get

$$|R_n - q^n| \leq \alpha_0,$$

i.e. the result for clean population without faulty elements (see [1]).

Example. Consider the inflated geometric distribution

$$p_n = \begin{cases} (1-\delta)a + \delta, & n=1, \\ (1-\delta)a(1-a)^{n-1}, & n \geq 2. \end{cases}$$

i.e. the distribution (p_n) given by (1), where

$$p'_n = a(1-a)^{n-1}, \quad p''_n = \begin{cases} 1, & n=1, \\ 0, & n=2,3,\dots, \end{cases}$$

and $\delta \in (0,1)$.

Then

$$R'_n = (1-a)^n, \quad R''_n = \begin{cases} 1, & n=0, \\ 0, & n=1,2,\dots, \end{cases}$$

$$\bar{G}'_n = (1-a)^{n+1}, \quad \bar{G}''_n = \begin{cases} 1, & n=0, \\ 0, & n=1,2,\dots, \end{cases}$$

and

$$\begin{aligned} \mu' &= \frac{1}{a}, & \mu'' &= 1, \\ \sigma^{2'} &= \frac{1-a}{a^2}, & \sigma^{2''} &= 0. \end{aligned}$$

Hence

$$R_n = \begin{cases} 1, & n=1, \\ (1-\delta)(1-a)^n, & n=1,2,\dots, \end{cases}$$

$$\bar{G}'_n = \begin{cases} (1-\delta)(1-a) + \delta, & n = 0, \\ (1-\delta)(1-a)^{n+1}, & n = 1, 2, \dots, \end{cases}$$

and

$$\mu = \frac{1-\delta+\delta a}{a}, \quad \sigma^2 = \frac{1-a-\delta+\delta a}{a^2}.$$

In this case we have by (4) and (9) that

$$\beta_\delta := \frac{(1+2\varepsilon)(1-\delta+\delta a)^2 - \delta a^2 + \delta - 1}{a(1+\varepsilon)(2-2\delta+2\delta a)}.$$

But by $\varepsilon = \frac{\delta(2-\delta)}{(1-\delta)^2}$, we have

$$\beta_\delta = \frac{(3+2a-a^2)\delta + (-7+2a+3a^2)\delta^2 + (5-6a+a^2)\delta^3 - (1-2a+a^2)\delta^4}{2a(1-\delta)^2 - (1-\delta+\delta a)},$$

omitting the terms with higher powers of δ we obtain the estimate

$$\beta_\delta \approx \frac{(3+2a-a^2)\delta}{2a(1-3\delta+\delta a)}.$$

For instance it $a = 0.9$ and $\delta = 0.01$ then the error of approximation $\beta_\delta = 0.0226$.

ACKNOWLEDGEMENTS: *I would like to thank to Professor Dobiesław Bobrowski for very many helpful comments and inspirational discussions.*

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Received on 13.02.1997 and, in revised form, on 10.12.1997.